

## Correlations in Asynchronous Markets

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## Outline

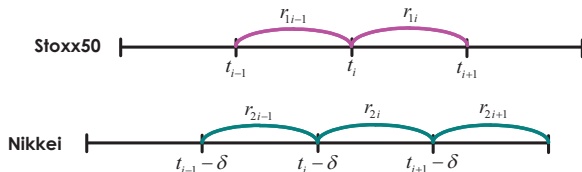
- Motivation
- Estimating correlations and volatilities in asynchronous markets
- Historical correlations: Stoxx50 – S&P500 – Nikkei
- Comparison with other heuristic estimators – options and correlation swaps
- Correlations larger than 1
- Conclusion

## Motivation

- Equity derivatives generally involve baskets of stocks/indices traded in different geographical areas
  - Operating hours of: Asian and European exchanges, Asian and American exchanges, usually have no overlap
  - Standard methodology on equity derivatives desks:
    - Use standard multi-asset model based on assumption of continuously traded securities
    - Compute/trade deltas at the close of each market, using *stale* values for securities not trading at that time
    - Likewise, valuation is done using *stale* values for securities whose markets are closed
- ▷ **How should we estimate volatility and correlation parameters ?**

## Correlation estimation in asynchronous markets

- Consider following situation:



- Valuation of the option is done at the close of the Stoxx50
  - Deltas are computed and traded on the market close of each security
- Daily P&L:

$$\begin{aligned}
 P\&L &= -[f(t_{i+1}, S_{1,i+1}, S_{2,i+1}) - f(t_i, S_{1,i}, S_{2,i})] \\
 &+ \frac{df}{dS_1}(t_i, S_{1,i}, S_{2,i})(S_{1,i+1} - S_{1,i}) \\
 &+ \frac{df}{dS_2}(t_i - \delta, S_{1,i-1}, S_{2,i})(S_{2,i+1} - S_{2,i})
 \end{aligned}$$

- Note that arguments of  $\frac{df}{dS_2}$  are different

- Rewrite delta on  $S_2$  so that arguments are same as  $f$  and  $\frac{df}{dS_1}$

$$\frac{df}{dS_2}(t_i - \delta, S_{1,i-1}, S_{2,i}) = \frac{df}{dS_2}(t_i, S_{1,i}, S_{2,i}) - \frac{d^2f}{dS_1 dS_2}(t_i, S_{1,i}, S_{2,i})(S_{1,i} - S_{1,i-1})$$

- Other correction terms contribute at higher order in  $\Delta$
- P&L now reads:

$$\begin{aligned} P\&L = & - [f(t_{i+1}, S_{1,i+1}, S_{2,i+1}) - f(t_i, S_{1,i}, S_{2,i})] \\ & + \frac{df}{dS_1}(S_{1,i+1} - S_{1,i}) + \left[ \frac{df}{dS_2} - \frac{d^2f}{dS_1 dS_2}(S_{1,i} - S_{1,i-1}) \right] (S_{2,i+1} - S_{2,i}) \end{aligned}$$

- Expanding at 2nd order in  $\delta S_1, \delta S_2$ :

$$P\&L = - \frac{df}{dt} \Delta - \left[ \frac{1}{2} \frac{d^2f}{dS_1^2} \delta S_{1+}^2 + \frac{1}{2} \frac{d^2f}{dS_1^2} \delta S_{2+}^2 + \frac{d^2f}{dS_1 dS_2} \delta S_{1+} \delta S_{2+} \right] - \frac{d^2f}{dS_1 dS_2} \delta S_{1-} \delta S_{2+}$$

- Assume  $f$  is given by a Black-Scholes equation:

$$\frac{df}{dt} + \frac{\sigma_1^2}{2} S_1^2 \frac{d^2 f}{dS_1^2} + \frac{\sigma_2^2}{2} S_2^2 \frac{d^2 f}{dS_2^2} + \rho \sigma_1 \sigma_2 S_1 S_2 \frac{d^2 f}{dS_1 dS_2} = 0$$

- P&L now reads:

$$\begin{aligned} P\&L = & -\frac{1}{2} S_1^2 \frac{d^2 f}{dS_1^2} \left[ \left( \frac{\delta S_1^+}{S_1} \right)^2 - \sigma_1^2 \Delta \right] - \frac{1}{2} S_2^2 \frac{d^2 f}{dS_2^2} \left[ \left( \frac{\delta S_2^+}{S_2} \right)^2 - \sigma_2^2 \Delta \right] \\ & - S_1 S_2 \frac{d^2 f}{dS_1 dS_2} \left[ \left( \frac{\delta S_1^-}{S_1} + \frac{\delta S_1^+}{S_1} \right) \frac{\delta S_2^+}{S_2} - \rho \sigma_1 \sigma_2 \Delta \right] \end{aligned}$$

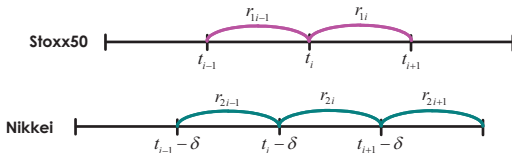
- Prescription for estimating volatilities & correlations so that P&L vanishes on average:

$$\sigma_1^{*2} = \frac{1}{\Delta} \left\langle \left( \frac{\delta S_1^+}{S_1} \right)^2 \right\rangle \quad \sigma_2^{*2} = \frac{1}{\Delta} \left\langle \left( \frac{\delta S_2^+}{S_2} \right)^2 \right\rangle \quad \rho^* \sigma_1^* \sigma_2^* = \frac{1}{\Delta} \left\langle \left( \frac{\delta S_1^-}{S_1} + \frac{\delta S_1^+}{S_1} \right) \frac{\delta S_2^+}{S_2} \right\rangle$$

- Volatility estimators are the usual ones, involving daily returns
- The correlation estimator involves daily returns as well

- Define  $r_i = \frac{\delta S_i^+}{S_i}$ . At lowest order in  $\Delta$ ,  $\frac{\delta S_i^-}{S_i} \simeq \frac{\delta S_{i-1}^-}{S_{i-1}}$

$$\sigma_1^{*2} = \frac{1}{\Delta} \langle r_{1i}^2 \rangle \quad \sigma_2^{*2} = \frac{1}{\Delta} \langle r_{2i}^2 \rangle \quad \rho^* = \frac{\langle (r_{1i-1} + r_{1i}) r_{2i} \rangle}{\sqrt{\langle r_{1i}^2 \rangle \langle r_{2i}^2 \rangle}}$$



- Had we chosen the close of the Nikkei for valuing the option: symmetrical estimator:

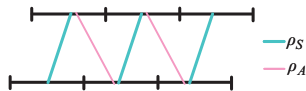
$$\sigma_1^{*2} = \frac{1}{\Delta} \langle r_{1i}^2 \rangle \quad \sigma_2^{*2} = \frac{1}{\Delta} \langle r_{2i}^2 \rangle \quad \rho^* = \frac{\langle r_{1i} (r_{2i} + r_{2i+1}) \rangle}{\sqrt{\langle r_{1i}^2 \rangle \langle r_{2i}^2 \rangle}}$$

- If returns are time-homogeneous  $\langle r_{1i-1} r_{2i} \rangle = \langle r_{1i} r_{2i+1} \rangle$   
 In practice  $\frac{1}{N} \sum_1^N (r_{1i-1} + r_{1i}) r_{2i} - r_{1i} (r_{2i} + r_{2i+1}) = \frac{1}{N} (r_{10} r_{21} - r_{1N} r_{2N+1})$

▷ Difference between two estimators of  $\rho^*$ : finite size effect of order  $\frac{1}{N}$

- In conclusion, in asynchronous markets: 2 correlations  $\rho_S, \rho_A$ :

$$\rho_S = \frac{\langle r_{1i} r_{2i} \rangle}{\sqrt{\langle r_{1i}^2 \rangle \langle r_{2i}^2 \rangle}} \quad \rho_A = \frac{\langle r_{1i} r_{2i+1} \rangle}{\sqrt{\langle r_{1i}^2 \rangle \langle r_{2i}^2 \rangle}}$$



and derivatives should be priced with  $\rho^*$ :

$$\rho^* = \rho_S + \rho_A$$

- ▷ Does  $\rho^*$  depend on the particular delta strategy used in derivation ?
- ▷ Is  $\rho^*$  in  $[-1, 1]$  ?
- ▷ How does  $\rho^*$  compare to standard correlations estimators evaluated with 3-day, 5-day,  $n$ -day returns ?



- What if had computed deltas differently – for example "predicting" the value of the stock not trading at the time of computation ?

- ▶ Option delta-hedged one way minus option delta-hedged the other way.  
 Final P&L is:

$$\sum (\Delta_t^a - \Delta_t^b)(S_{t+\Delta} - S_t)$$

- ▶ Price of pure delta strategy is zero: **correlation estimator is independent on delta strategy used in derivation**

- Imagine processes are continuous yet observations are asynchronous:  
 assume that  $\rho\sigma_1\sigma_2$ ,  $\sigma_1^2$ ,  $\sigma_2^2$  are periodic functions with period  $\Delta = 1$  day:

$$\rho_S = \frac{\frac{1}{\Delta} \int_t^{t+\Delta-\delta} \rho\sigma_1\sigma_2 ds}{\sqrt{\frac{1}{\Delta} \int_t^{t+\Delta} \sigma_1^2 ds} \sqrt{\frac{1}{\Delta} \int_{t-\delta}^{t+\Delta-\delta} \sigma_2^2 ds}}$$

$$\rho_A = \frac{\frac{1}{\Delta} \int_{t+\Delta-\delta}^{t+\Delta} \rho\sigma_1\sigma_2 ds}{\sqrt{\frac{1}{\Delta} \int_t^{t+\Delta} \sigma_1^2 ds} \sqrt{\frac{1}{\Delta} \int_{t-\delta}^{t+\Delta-\delta} \sigma_2^2 ds}}$$

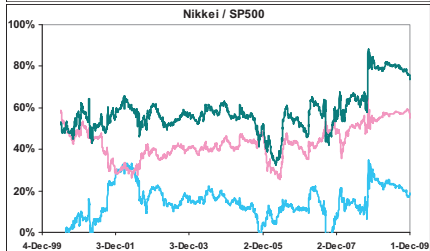
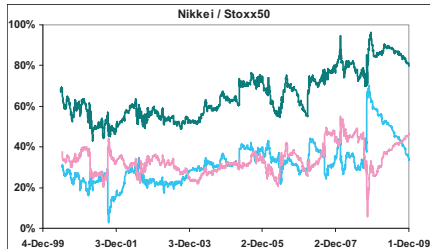
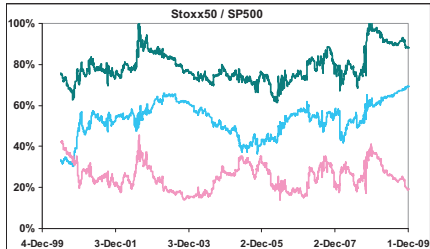
$$\rho^* = \rho_S + \rho_A$$



- ▶ Recovers value of "synchronous correlation": **no bias**

## Historical correlations

- $\rho_S$  (blue),  $\rho_A$  (pink),  $\rho^* = \rho_S + \rho_A$  (green) – 6-month EWMA



- $\rho_S, \rho_A$  seem to move antithetically
  - Imagine  $\sigma_1(s) = \sigma_1 \lambda(s)$ ,  $\sigma_2(s) = \sigma_2 \lambda(s)$ ,  $\rho$  constant, with  $\lambda(s)$  such that  $\frac{1}{\Delta} \int_0^\Delta \lambda^2(s) ds = 1$ . Then:

$$\rho_S = \rho \frac{1}{\Delta} \int_0^{\Delta-\delta} \lambda^2(s) ds$$

$$\rho_A = \rho \frac{1}{\Delta} \int_{\Delta-\delta}^\Delta \lambda^2(s) ds$$

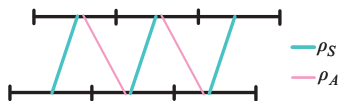
and  $\rho^*$  is given by:

$$\rho^* = \rho_S + \rho_A = \rho$$

- By changing  $\lambda(s)$  we can change  $\rho_S, \rho_A$ , while  $\rho^*$  stays fixed.
- ▷ The relative sizes of  $\rho_S, \rho_A$  are given by the intra-day distribution of the realized covariance.

## Comparison with heuristic estimators

- Trading desks have long ago realized that merely using  $\rho_S$  is inadequate
  - Standard fix: compute standard correlation using 3-day, 5-day, you-name-it, rather than daily returns
  - How do these estimators differ from  $\rho^*$  ?
- Connected issue: how do we price an  $n$ -day correlation swap ?



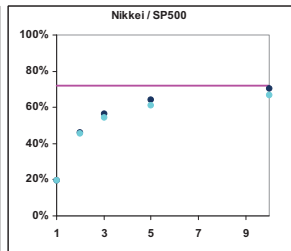
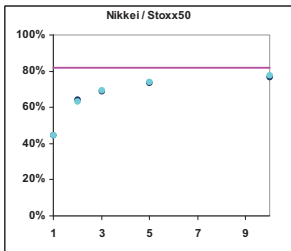
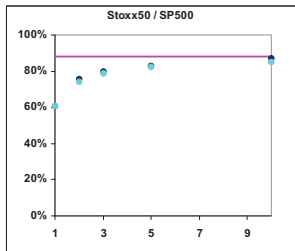
- ▷ An  $n$ -day correlation swap should be priced with  $\rho_n$  given by:

$$\rho_n = \rho_S + \frac{n-1}{n} \rho_A$$

- For  $n = 3$ ,  $\rho_3 = \rho_S + \frac{2}{3} \rho_A$
- If no serial correlation in historical sample, standard correlation estimator applied to  $n$ -day returns yields  $\rho_n$

## Historical $n$ -day correlations

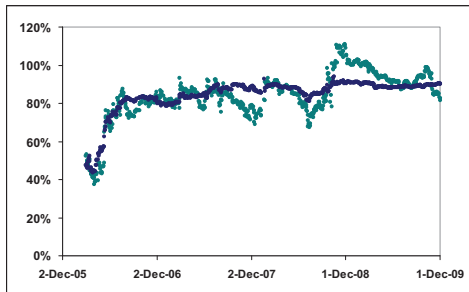
- $n$ -day correlations evaluated on 2004-2009 with:
  - $n$ -day returns (dark blue)
  - using  $\rho_S + \frac{n-1}{n}\rho_A$  (light blue)
- compared to  $\rho^*$  (purple line)



- Common estimators  $\rho_3, \rho_5$  underestimate  $\rho^*$

## The S&P500 and Stoxx50 as synchronous securities

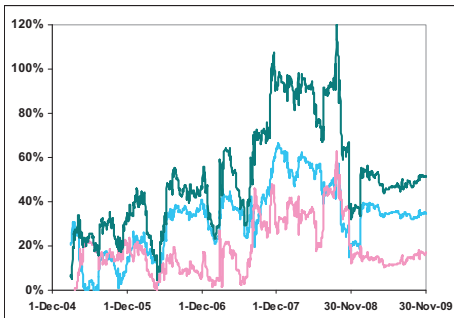
- European and American exchanges have some overlap. We can either:
  - delta-hedge asynchronously the S&P500 at 4pm New York time and the Stoxx50 at 5:30pm Paris time
  - delta-hedge simultaneously both futures at – say – 4pm Paris time
- 1st case: use  $\rho^*$ , 2nd case: use standard correlation for synchronous securities – are they different ?
- $\rho^*$  (light blue), standard sync. correlation (dark blue) – 3-month EWMA



- Matches well, but not identical: difference stems from residual *realized* serial correlations.

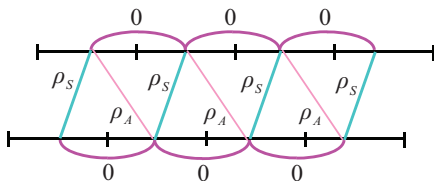
## Correlations larger than 1

- Example of RBS/Citigroup correlations:  $\rho_S$  (blue),  $\rho_A$  (pink),  $\rho^*$  (green) – 3-month EWMA



- Are instances when  $\rho^* > 1$  an artifact ? Do they have financial significance ?

- Consider a situation when no serial correlation is present. The global correlation matrix is positive, by construction. How large can  $\rho_S + \rho_A$  be ?



- Compute eigenvalues of full correlation matrix:
  - assume both ladder uprights consist of  $N$  segments, with periodic boundary conditions
  - assume eigenvalues have components  $e^{ik\theta}$  on higher upright,  $\alpha e^{-ik\theta}$  on lower upright
  - express that  $\lambda$  is an eigenvalue:

$$\begin{aligned} \alpha \rho_S + 1 + \alpha e^{i\theta} &= \lambda \\ \rho_S + \alpha + e^{-i\theta} \rho_A &= \lambda \alpha \end{aligned}$$

yields:

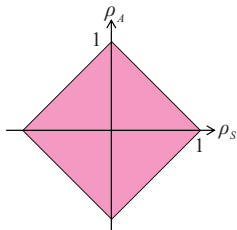
$$\lambda = 1 \pm \sqrt{(\rho_S + \rho_A \cos \theta)^2 + \rho_A^2 \sin^2 \theta}$$



- Periodic boundary conditions impose  $\theta = \frac{2n\pi}{N}$ , where  $n = 0 \dots N - 1$
- $\lambda(\theta)$  extremal for  $\theta = 0, \pi$ . For these values  $\lambda = 1 \pm |\rho_S \pm \rho_A|$
- $\lambda > 0$  implies:

$$-1 \leq \rho_S + \rho_A \leq 1$$

$$-1 \leq \rho_S - \rho_A \leq 1$$



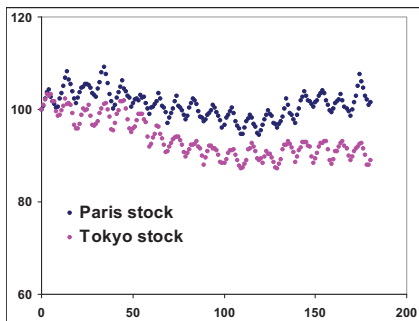
- ▷ **If no serial correlations  $\rho^* \in [-1, 1]$**
- ▷ Instances when  $\rho^* > 1$ : evidence of serial correlations
- ▷ **Impact of  $\rho^* > 1$  on trading desk: price with the right realized volatilities, 100% correlation  $\rightarrow$  lose money !!**

## Example with basket option

- Sell 6-month basket option on basket of Japanese stock & French stock.

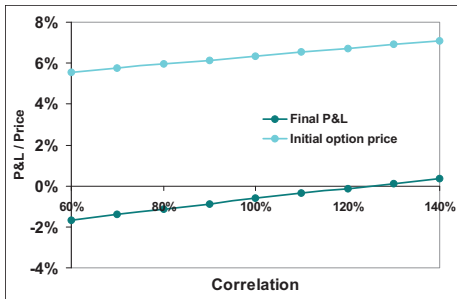
$$\text{Payoff is } \left( \sqrt{\frac{S_1^T S_2^T}{S_1^0 S_2^0}} - 1 \right)^+$$

- Basket is lognormal with volatility given by  $\sigma = \sqrt{\sigma_1^2 + \sigma_2^2 + 2\rho\sigma_1\sigma_2}$
- Use following "historical" data:



- Realized vols are 21.8% for  $S_1$ , 23.6% for  $S_2$ . Realized correlations are  $\rho_S = 63.3\%$ ,  $\rho_A = 57.6\%$ :  $\rho^* = 121\%$ .

- Backtest delta-hedging of option with:
  - implied vols = realized vols
  - different implied correlations
- Initial option price and final P&L:



- Final P&L vanishes when one prices and risk-manages option with an implied correlation  $\rho \approx 125\%$ .

## Conclusion

- It is possible to price and risk-manage options on asynchronous securities using the standard synchronous framework, provided special correlation estimator is used.
- Correlation estimator quantifies correlation that is materialized as cross-Gamma P&L.
- Same technique applies to asynchronous FX / Swap rates / ...
- Correlation swaps and options have to be priced with different correlations.
- Serial correlations may push realized value of  $\rho^*$  above 1: a short correlation position will lose money, even though one uses the right vols and 100% correlation.